## SUBSPACE MODELS

## Credits

$\square$ Some of these slides were sourced and/or modified from Simon Prince, University College London

## Subspace Models

$\square$ Natural images have high dimensionality $D$

- e.g., for an $1800 \times 1200$ colour image, $D \cong 6.5$ million.
$\square$ There is typically insufficient training data to learn a probabilistic model in such a high-dimensional space.
$\square$ Fortunately, natural images actually live in a much smaller subspace, or manifold, of this highdimensional space.



## Subspace Models

$\square$ For example, you will have to wait a long time before a sample of white noise looks like a natural image.


## Subspace Models

$\square$ e.g., standard transformations (e.g., translations, rotations, scalings) of objects produce images populating a low-dimensional manifold embedded in this high-dimensional space


## Subspace Models

$\square$ The goal of subspace methods is to discover the low-dimensional subspace in which the data lie and exploit the lower-dimensionality to allow efficient and detailed modeling.

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## Subspace Models

$\square$ We will mainly consider linear subspaces
$\square$ A line if $D=2$
$\square$ A line or a plane if $D=3$
$\square$ A hyperplane of dimensionality [ $1, \ldots, D-1]$ for higher $D$
$\square$ But we will also consider some methods to deal with nonlinear manifolds.

## PRINCIPAL COMPONENT ANALYSIS

## Principal Component Analysis

$\square$ PCA finds the linear subspace that
$\square$ maximizes the explained variance
$\square$ equivalently, minimizes the unexplained variance
$\square$ PCA can be applied to any multidimensional dataset

- (data do not have to be Gaussian)



## Maximum Variance Formulation

Observations $\left\{\mathbf{x}_{\mathrm{n}}\right\}, n=1, \ldots N$
Observation $\mathbf{x}_{\mathrm{n}}$ is a high-dimensional vector of dimension $D$
Let $\overline{\mathbf{x}}=\frac{1}{N} \sum_{i=1}^{N} \mathbf{x}_{\mathrm{n}}$ be the sample mean and $\mathbf{S}=\frac{1}{N} \sum_{i=1}^{N}\left(\mathbf{x}_{\mathrm{n}}-\overline{\mathbf{x}}\right)\left(\mathbf{x}_{\mathrm{n}}-\overline{\mathbf{x}}\right)^{t}$ be the sample covariance
Goal: Project the data onto subspace of dimension $M<D$

Consider a direction in the data space given by unit vector $u_{1}$.

Now imagine projecting all of the data onto this unit vector.

The mean of the projected data is $\mathbf{u}_{1}^{\mathrm{t}} \overline{\mathbf{x}}$.
The variance of the projected data is $\frac{1}{N} \sum_{i=1}^{N}\left(\mathbf{u}_{1}^{\mathrm{t}} \mathbf{x}_{\mathrm{n}}-\mathbf{u}_{1}^{\mathrm{t}} \overline{\mathbf{x}}\right)^{2}=\mathbf{u}_{\mathbf{1}}^{\mathrm{t}} \mathbf{S} \mathbf{u}_{1}$


## Maximum Variance Formulation

## Subspace Models

We want to select the unit vector $u_{1}$ that maximizes the projected variance $u_{1}^{\mathrm{t}} S \mathrm{u}_{1}$

To do this, we use a Lagrange multiplier $\lambda_{1}$ to maintain the constraint that $\mathbf{u}_{1}$ be a unit vector.

Thus we seek to maximize $\mathbf{u}_{1}^{\mathrm{t}} \mathbf{S u}_{1}+\lambda_{1}\left(1-\mathbf{u}_{1}^{\mathrm{t}} \mathbf{u}_{1}\right)$

Setting the derivative with respect to $u_{1}$ to 0 , we have $S u_{1}=\lambda_{1} u_{1}$

Thus $\mathbf{u}_{1}$ is an eigenvector of $\mathbf{S}$.

Left-multiplying by $\mathbf{u}_{1}^{\mathrm{t}}$, we see that the projected variance $\mathbf{u}_{1}^{\mathrm{t}} \boldsymbol{S} \mathbf{u}_{1}=\lambda_{1}$.


## Dimensionality Reduction

$\square$ The next direction $\mathbf{U}_{\mathbf{2}}$ can be chosen by maximizing projected variance in the $D-1$ dimensional subspace orthogonal to $\mathbf{u}_{1}$.
$\square$ Typically, most of the variance is captured in a relatively small linear subspace.


## Computational Cost

$\square$ Computing full eigenvector decomposition is $O\left(D^{3}\right)$.
$\square$ If we only need the first $M$ eigenvectors, the cost is $O\left(M D^{2}\right)$.
$\square$ However, this could still be very expensive if $D$ is large
e.g., For an $1800 \times 1600$ image and $M=100, O(650$ million $)$

## Computational Cost

$\square$ But the number of training images $N$ is usually much smaller than $D$, and this leads to a trick:

Let $\mathbf{X}$ be the $N \times D$ centred data matrix whose $n$th row is given by $\left(\mathbf{x}_{\mathrm{n}}-\overline{\mathbf{x}}\right)^{t}$.
Then the sample covariance matrix is $\mathbf{S}=\frac{1}{N} \mathbf{X}^{t} \mathbf{X}$.
and the eigenvector equation is $\frac{1}{N} \stackrel{D \times D}{\mathbf{X}^{t}} \mathbf{X}_{\mathbf{i}}=\lambda_{i} \mathbf{u}_{i}$
Pre-multiplying both sides by $\mathbf{X}$ yields $\frac{1}{N} \mathbf{X} \mathbf{X}^{\mathbf{t}}\left(\mathbf{X u}_{\mathbf{i}}\right)=\lambda_{i}\left(\mathbf{X} \mathbf{u}_{\mathbf{i}}\right)$
Now letting $\mathbf{v}_{\mathbf{i}}=\mathbf{X} \mathbf{u}_{\mathbf{i}}$, we have
$\frac{1}{N} \stackrel{N \times N}{\stackrel{N}{X}}{ }^{t} \mathbf{v}_{\mathbf{i}}=\lambda_{i} \mathbf{v}_{\mathbf{i}}$

## Computational Cost

$\square$ To find the eigenvectors of $\mathbf{S}$, we premultiply by $\mathbf{X}^{\text {t: }}$
$\frac{1}{N} \stackrel{N \times N}{\mathbf{X X}^{t} \mathbf{v}_{\mathbf{i}}}=\lambda_{i} \mathbf{v}_{\mathbf{i}} \rightarrow \overbrace{\left(\frac{1}{N} \mathbf{X}^{\mathbf{t}} \mathbf{X}\right)}^{S}\left(\mathbf{X}^{\left.\mathbf{t} \mathbf{v}_{\mathbf{i}}\right)}=\lambda_{i}\left(\mathbf{X}^{\mathbf{t}} \mathbf{v}_{\mathbf{i}}\right)\right.$
and, normalized to unit length, the eigenvectors are $\mathbf{u}_{\mathbf{i}}=\frac{1}{\sqrt{N \lambda}} \mathbf{X}^{\mathbf{t}} \mathbf{v}_{\mathbf{i}}$
Note that these $N$ eigenvectors live in the $N$-dimensional subspace spanned by the training images.

## Pre-Whitening



Original Data


Normalized to 0 -mean and unit variance (z-scores)


Whitened

## Compression



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## Modeling



Low-dimensional model of variation of registered objects such as faces

$\overline{\mathrm{x}}+{ }_{1} q \mathbf{u}_{1}$

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## PROBABILISTIC PCA

## Probabilistic PCA

$\square$ PCA applies to data of any kind.
$\square$ But PCA can also be interpreted as the maximum likelihood solution to a probabilistic latent variable model based on a constrained form of the Gaussian distribution:

Let $\mathbf{z}$ be an M-dimensional hidden variable with Gaussian prior
Let $\mathbf{x}$ be the D-dimensional observed variable with Gaussian conditional:
$p(\mathbf{z})=N(\mathbf{z} \mid 0, \mathbf{I})$
$p(\mathbf{x} \mid \mathbf{z})=N\left(\mathbf{x} \mid \mathbf{W} \mathbf{z}+\mu, \sigma^{2} \mathbf{I}\right)$


## Probabilistic PCA

$\square$ One way to see this is to think of Probabilistic PCA as the limit of a mixture of Gaussians model, as the number of Gaussian components $\rightarrow \infty$ :

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## Probabilistic PCA

Consider putting the means of the Gaussians mixture components all on a line and forcing their diagonal covariances to be identical.
What happens if we keep adding more and more Gaussians along this line?


J. Elder

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## Probabilistic PCA

$\square$ PCA as the maximum likelihood solution to a probabilistic latent variable model based on a constrained form of the Gaussian distribution:

Let $\mathbf{z}$ be an M-dimensional hidden variable with Gaussian prior
Let $\mathbf{x}$ be the D-dimensional observed variable with Gaussian conditional:
$p(\mathbf{z})=N(\mathbf{z} \mid 0, \mathbf{I})$
$p(\mathbf{x} \mid \mathbf{z})=N\left(\mathbf{x} \mid \mathbf{W} \mathbf{z}+\mu, \sigma^{2} \mathbf{I}\right)$
$\uparrow_{D \times M}$

## Generative Model



## Marginal Distribution for Probabilistic PCA

$\square$ The marginal distribution of the observed variable is
$p(x)=N(\mathbf{x} \mid \mu, \mathbf{C})$
where
$\mathbf{C}=\mathbf{W W}^{t}+\sigma^{2} \mathbf{I}$

## Maximum Likelihood PCA

$\mu=\overline{\mathbf{x}}$
$\sigma_{M L}^{2}=\frac{1}{D-M} \sum_{i=M+1}^{D} \lambda_{i}$
$\mathbf{W}_{M L}=\mathbf{U}_{M}\left(\mathbf{L}_{M}-\sigma^{2} \mathbf{I}\right)^{1 / 2} \mathbf{R}$
where
$\mathbf{U}_{M}$ is a $D \times M$ matrix whose columns are given by any subset of size $M$ of the eigenvectors of $\mathbf{S}$
$\mathbf{L}_{M}$ is an $M \times M$ diagonal matrix containing the $M$ corresponding eigenvalues $\lambda_{i}$
$\mathbf{R}$ is an arbitrary rotation matrix
ML parameters can be found either by
$\square$ determining $M$ eigenvectors and eigenvalues directly
$\square \mathrm{EM} \quad \mathbf{x}=\mathbf{W z}+\mu+\varepsilon$

## FACTOR ANALYSIS

## Generative Model



## Factor Analysis Terminology

$\square$ Columns of W are called factor loadings
$\square$ Diagonal elements of $\Psi$ are called uniquenesses


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## Marginal Distribution for Factor Analysis

$\square$ The marginal distribution of the observed variable is
$p(x)=N(\mathbf{x} \mid \mu, \mathbf{C})$
where
$\mathbf{C}=\mathbf{W W}^{t}+\Psi$

## Maximum Likelihood Parameter Estimation

$$
\mu_{M L}=\bar{x}
$$

However, no closed-form solution for $\mathbf{W}_{M L}$. Instead, can use EM to find $\mathbf{W}_{M L}$ and $\Psi_{M L}$.

## Learning Results: Two Factor Model

## Subspace Models

$\mathbf{X}=\mathbf{W} \mathbf{z}+\mu+\boldsymbol{\varepsilon}$ where $\mathbf{W}=\left[\mathbf{w}_{1} \mathbf{w}_{2} \ldots \mathbf{w}_{D}\right]$

$\mu$

$\mathbf{W}_{1}$


$$
\mu+2 \mathbf{w}_{1} \quad \mu+2 \mathbf{w}_{2}
$$



## Learning results: Two Factor Model

## Subspace Models

$\mathbf{X}=\mathbf{W} \mathbf{z}+\mu+\boldsymbol{\varepsilon}$ where $\mathbf{W}=\left[\mathbf{w}_{1} \mathbf{w}_{2} \ldots \mathbf{w}_{D}\right]$

$\mu$

$\mathbf{W}_{1}$


$$
\mu+2 \mathbf{w}_{1} \quad \mu+2 \mathbf{w}_{2}
$$


$\Psi$

## Non-Linear Extensions

Mixture of factor analyzers (MOFA)
$\square$ Two levels of the EM algorithm
$\square$ One to learn each factor analyzer

- One to learn the mixture model
$\square$ Can describe quite complex manifold structures in high dimensions with only a limited number of parameters


$$
p(x)=\sum_{k=1}^{K} N\left(\mathbf{x} \mid \mu_{k}, \mathbf{C}_{k}\right)
$$

where

$$
\mathbf{C}_{k}=\mathbf{W}_{k} \mathbf{W}_{k}^{t}+\Psi_{k}
$$

Pixel 1

## Non-Linear Extensions

$\square$ Kernel PCA

- Idea:
- Use a non-linear mapping $\phi$ to an $M$-dimensional 'feature space'
- Now perform PCA in this new space




## Kernel PCA

Assume 0-mean data vectors: $\sum_{n=1}^{N} \mathbf{x}_{n}=0$.
M-dimensional eigenvector
Covariance in feature space $C=\frac{1}{N} \sum_{n=1}^{N} \phi\left(x_{n}\right) \phi\left(x_{n}\right)^{t}$
Eigenvector expansion $C v_{i}=\lambda v_{i}$

$$
] \rightarrow \frac{1}{N} \sum_{n=1}^{N} \phi\left(x_{n}\right)\left\{\phi\left(x_{n}\right)^{t} v_{i}\right\}=\lambda_{i} v_{i}
$$

Thus the eigenvector $v_{i}$ is a linear combination of the transformed data vectors $\phi\left(x_{n}\right)$ :
$v_{i}=\sum_{n=1}^{N} a_{i n} \phi\left(x_{n}\right)$
Substituting, we have $\frac{1}{N} \sum_{n=1}^{N} \phi\left(x_{n}\right) \phi\left(x_{n}\right)^{t} \sum_{m=1}^{N} a_{i m} \phi\left(x_{m}\right)=\lambda \sum_{n=1}^{N} a_{i n} \phi\left(x_{n}\right)$
Now multiplying both sides by $\phi\left(x_{l}\right)^{t}$, we obtain
$\frac{1}{N} \sum_{n=1}^{N} \phi\left(x_{l}\right)^{t} \phi\left(x_{n}\right) \sum_{m=1}^{N} a_{i m} \phi\left(x_{n}\right)^{t} \phi\left(x_{m}\right)=\lambda_{i} \sum_{n=1}^{N} a_{i n} \phi\left(x_{l}\right)^{t} \phi\left(x_{n}\right)$
Finally, defining the kernel function $\mathrm{k}\left(\mathrm{x}_{n}, x_{m}\right)=\phi\left(x_{n}\right)^{t} \phi\left(x_{m}\right)$, we can write $\frac{1}{N} \sum_{n=1}^{N} k\left(x_{1}, x_{n}\right) \sum_{m=1}^{N} a_{i m} k\left(x_{n}, x_{m}\right)=\lambda_{i} \sum_{n=1}^{N} a_{i n} k\left(x_{1}, x_{n}\right)$

## Kernel PCA

$\frac{1}{N} \sum_{n=1}^{N} k\left(x_{l}, x_{n}\right) \sum_{m=1}^{N} a_{i m} k\left(x_{n}, x_{m}\right)=\lambda_{i} \sum_{n=1}^{N} a_{i n} k\left(x_{l}, x_{n}\right)$
N -dimensional eigenvector
or, in matrix notation, $\mathbf{K}^{2} \mathbf{a}_{i}=\lambda_{i} N \mathbf{K a}_{i} \rightarrow \mathbf{K a}_{i}=\lambda_{i} N \mathbf{a}_{i}$

Requiring that the eigenvectors $\mathbf{v}_{i}$ in feature space be unit vectors leads to the constraint $\left|\mathbf{a}_{i}\right|^{2}=\frac{1}{N \lambda_{i}}$.

In practice, for the projected data to have 0 mean, use $\tilde{K}=K-\mathbf{1}_{N} K-K 1_{N}+\mathbf{1}_{N} K 1_{N}$ where

$$
\mathbf{1}_{N}=\underset{N}{\left(\begin{array}{ccc}
\frac{1}{N} & \cdots & \frac{1}{N} \\
\vdots & \ddots & \vdots \\
\frac{1}{N} & \cdots & \frac{1}{N}
\end{array}\right)} \downarrow N N
$$

## Kernel PCA

## Subspace Models

Eigenvalue $=21.72$


Eigenvalue=3.66


Eigenvalue $=21.65$


Eigenvalue $=3.09$


Eigenvalue=4.11


Eigenvalue=2.60


Eigenvalue=3.93


Eigenvalue=2.53


